

Full Length Research

An econometric analysis: Is there an optimal external debt threshold for Guinea?

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This study aimed to determine the threshold level of public external debt-to-GDP ratio for Guinea and compared it with the current evolution of public debt. The authors used an autoregressive distributed lag (ARDL) approach to test the hypothesis of long-term relationship among the variables of interest and the data used ranged from 1990 to 2018. The results revealed that External debt-to-GDP ratio and per capita GDP are positively related. Moreover, the null hypothesis of no co-integration was rejected. Thus, external debt-to-GDP ratio and economic growth as well as other variables are co-integrated. Moreover, the debt variable had significant non-linear effects on economic growth and indicated that there exists an optimal level of external debt-to-GDP ratio that stood at 25.2%. Compared to its current level which stood at 21.7%, the country still has some borrowing margin. In the short run external debt-to-GDP ratio has no significant effect on the country's economic performance.

Key words: External debt, Economic growth, autoregressive distributed lag (ARDL), Guinea.

INTRODUCTION

The problem of external public debt emerged in most developing countries during the 1970s and in the early 1980s for Africa. With an external debt of \$337.2 billion (1999), Africa is the most indebted continent in relation to its gross national product (GNP). This situation can be explained by the ease with which African countries were able to borrow money in the late 1970s and early 1980s, but also by the duration of the global economic crisis and the fall in the prices of raw materials and agricultural products. In addition, external debt has often impacted

the economy of over-indebted countries by taking away from the state budget resources necessary for the proper functioning of public administrations and services thereby reducing the country's capacity for investment, etc. It can also lead to an increased need for additional resources and hence a need for more borrowing. Faced with macroeconomic and financial imbalances in their economies, governments have often resorted to external borrowing to rebalance their external and internal deficits. The idea that debt servicing negatively affects the

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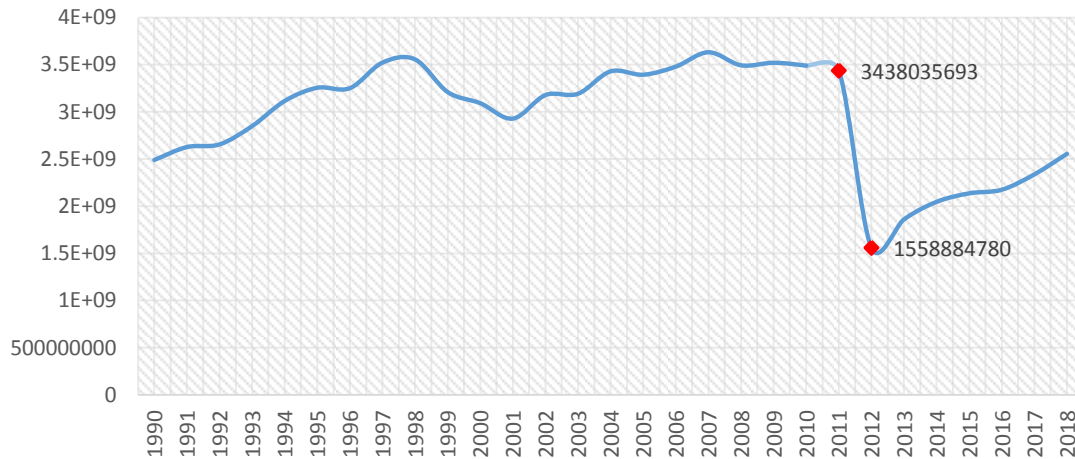


Figure 1. Evolution of the external debt stock (in current \$).
Source: World Bank (2020).

economy of poor countries has been the subject of sustained discussions within international institutions. This could explain why numerous programs aimed at relieving debt were put in place that is, the HIPC (Highly Indebted Poor Country) initiative. Indeed, many countries, including Guinea, were facing economic difficulties due to the large share of their GDP that they were allocating to debt service. Thus, the country embarked on a borrowing spree both internally and externally (International Monetary Fund, 2017). These borrowings undoubtedly affected the country's public finances.

Indeed, between the year 2000 and 2005, Guinea's external debt stock rose from US\$ 3 billion to US\$ 3.4 billion (World Bank, 2020). In 2011, the country's external debt stood at US\$ 3.4 billion. It dropped to US\$ 1.56 billion in 2012 in line with the HIPC initiative, representing a 54.66% reduction of the country's external debt. Notwithstanding this, the country's external debt kept rising steadily. By 2018, the country's external debt has reached another peak. Indeed, it stood at US 2.55 billion (World Bank, 2020). This represented a 63.88% increase in a six years period. The current observation is that Guinea's external debt is still on an upward sloping trend. Thus, in order to avoid falling back into the cycle of debt overhang with its adverse effects on both economic growth and fiscal sustainability, the mobilization of foreign resources should be carried out in a prudent manner, taking into account the costs and risks associated with these resources. It is in line with the above that, this study seeks to provide a better understanding of the following issues:

1. How did the country's public debt and its components evolved from 1990 to 2018?
2. What could be the short- and long-term impact of the country's rising external debt on economic growth?
3. Whether there is a critical threshold for Guinea's external debt?

The overall objective of this study is therefore to provide a better understanding of the impact of external debt on Guinea's economy. More specifically, the study seeks to: determine the short and long term impact of the country's external debt on its economic growth; determine the optimal threshold for Guinea's external debt. In addition to the above objectives, the following hypotheses will be tested: external debt has a positive impact on Guinea's economic growth in both the short and long run; and Guinea's optimal external debt-to-GDP ratio is below the current level of external debt which stood at 21.7%.

Stylized facts

As countries seek ways to ensure macroeconomic stability and mobilize resources for the financing of major development projects they resort to debt. However, in many instances, public debt has a number of limitations that should not be overlooked. Indeed, poor management of public debt can lead to a number of difficulties that can be very unfavorable to the proper functioning of the economy. In Guinea, the increase in external debt in recent years could be explained by the implementation of several major economic and social development projects. Thus, a descriptive analysis of external debt, as well as its relationship with certain key macroeconomic variables, should provide a better understanding. From Figure 1, Guinea's external debt increased from US\$ 2.489 billion to US\$ 3.55 billion between 1990 and 1998, an increase of 42.82%. However, over the period of 1998 to 2008, Guinea's external debt showed a moderate evolution, with light fluctuations.

The year 2012 was marked by a sharp drop of the country's external debt, this was mainly attributable to the cancellation of debt resulting from the implementation of the Heavily Indebted Poor Countries Initiative (HIPC), that is, a cancellation of more than US\$ 2 billion. From 2012

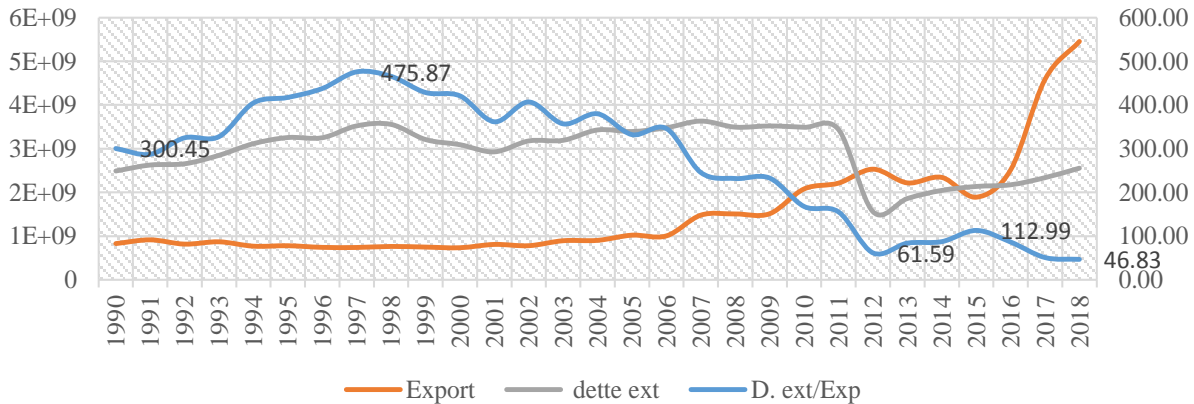


Figure 2. Evolution of the external debt to exports of goods and services ratio.
Source: World Bank, (2020).

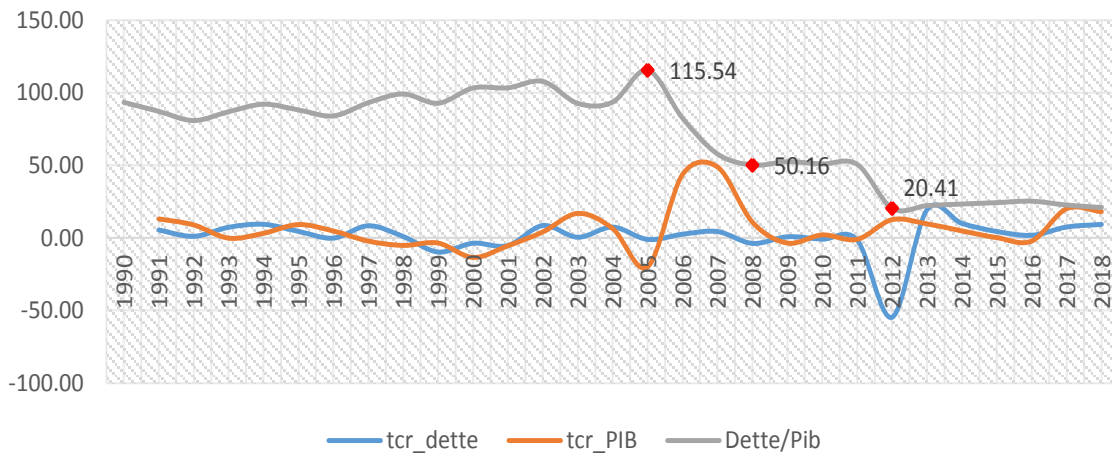


Figure 3. Trend of the growth rate of debt, GDP and debt-to-GDP ratio from 1990 to 2018.
Source: World Bank (2020).

to 2018, the stock of external debt rose from USD 1.558 billion to USD 2.554 billion, an increase of 63.88%. This increase can be explained by the State's investments in the mining and infrastructure sectors¹. When considering the external debt-to – export ratio, it is observed that the trend was upward sloping in the early 90s (Figure 2). Indeed, it rose from 300.45% in 1990 to 475.87% in 1997 representing a 58.36% rise. From 1997, the trend has been downward sloping till 2012 although not steadily. In 2012 the external debt-to- exports ratio stood at 61.59%. Unfortunately this downward trend was halted in the period ranging from 2013 to 2015 where it reached 112.99%. Thereafter, it dropped to reach its lowest level in 2018 where it stood at 46.83% representing a 58.55% decrease in less than five years (World Bank, 2020).

In Figure 3, it is observed that the growth rate of the economy and that of the external debt evolved together but with different amplitudes. Indeed, between 1990 and 2018, the outstanding external debt has grown by an average of 1.17%. As for the economic growth rate, it has evolved on average by 6.45% over the same period. Over the period 1990 to 2004, the external debt/GDP ratio was almost stable with an average growth rate of 2.61%. From 2005 to 2012, there was a clear decline in the external debt to GDP ratio. Indeed, it fell from 115.54% in 2005 to 50.16 in 2008 and then to its lowest level in 2012 at 20.41%. This later one has to do with the HIPC initiative. From 2012, the external debt to GDP ratio evolved around 21% (World Bank, 2020).

The trend of the country's debt service is presented in Figure 4. It is observed that it declined until 1992. Then it rose from 1993 to 1995 before another decline in 1996. Broadly speaking was neither steady nor stable

¹ Le Programme d'Investissement Public / Guinée 2019 (PIP)

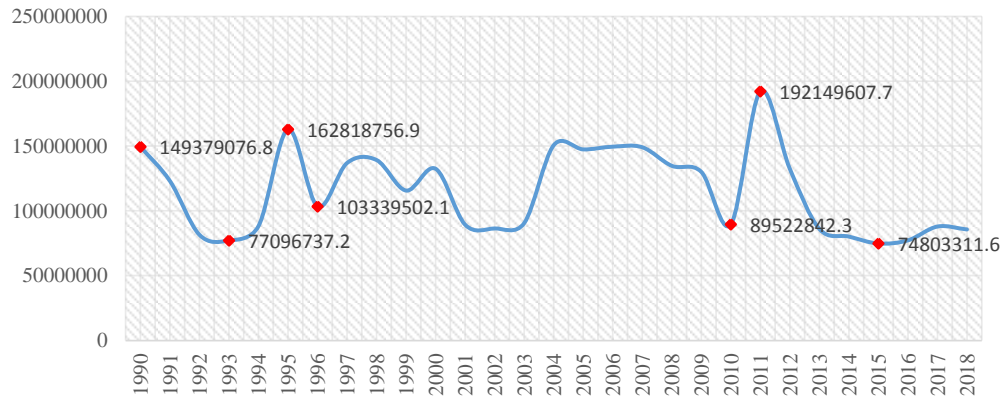


Figure 4. Evolution of debt service.
Source: World Bank, (2020).

throughout the period of analysis. However, from 2011, a downward trend is observed. This last situation is attributable to the cancellation of a large part of the external debt through the Heavily Indebted Poor Countries Initiative (HIPC).

REVIEW OF LITERATURE

A review of selected literature is undertaken in this study. It starts with a brief theoretical review followed by a selected empirical studies review on the link between public debt and economic growth.

Theoretical review

The debate between economic growth and debt is relatively old, one of the pioneers being Cairnes (1874), and owes its revival to endogenous growth theories. Since the 1980s, two schools of thoughts have been clashing over the theory of growth and public debt, namely the Keynesians and the neoclassicals. For the Keynesians, the main idea is that debt does not cause burdens for current and future generations, because of the investments it generates. From this approach, debt stimulates demand, and the accelerating effect of an increase in investment leads to an increase in production. According to Clements et al. (2004), external debt has the potential to stimulate economic growth, provided that it is used to finance investment. For these authors, it is necessary to have a measure of debt, because there is a certain threshold beyond which debt negatively influences economic growth. Indeed, when the return on capital is declining, the benefits of any new investment on economic growth could diminish as the debt increases. This theory gives rise to a "Laffer Curve" relationship between external debt on the one hand and per capita

income growth on the other. For the neoclassicals, debt is considered as a future tax and attributes it to the state. According to this school of thought, public debt has a negative effect on the accumulation of capital and the consumption of future and present generations. According to Sargent (1981), a sustainable debt leads to a growth rate higher than the real interest rate on bonds. Thus, government revenues grow faster than interest on the debt, based on the assumption of a unitary elasticity between the budget balance and economic activity. Krugman (1988) and Sachs (1989) predict that a high debt is harmful to economic growth, since it discourages investment. For these authors, when debt exceeds a country's internal resources, the country may no longer be able to repay past loans, which will have a dissuasive effect on potential creditors and investors thus, it hinders economic growth. Moreover, Barro's (1990) model attributed a very important role to productive public spending (for instance public spending on infrastructure) in the process of long-run economic growth. According to the author, debt is neither a wealth for the current generation or a bridge between generations because of the agents' anticipation of future taxes. Thus, part of the debt will be transferred to the future generation (tax debt) and the other part will be compensated by public securities. This is why substituting borrowing for taxation does not necessarily lead to growth. In matters of fiscal policy, public debt is a key factor in analyzing government's room for maneuver in its spending.

Empirical review

Several studies have tackled the issue of the optimal level of external debt to GDP ratio. However, there is no consensus on the threshold to be considered. This could be explained by the peculiarities of countries and their specific needs. This study therefore reviews couple of selected empirical studies including Greenidge et al.

(2012), Mencinger et al. (2014), Faye and Thiam (2015), Wade (2015), Omotosho et al. (2016), Adeniyi et al. (2018), N'Zué (2018), Mary et al. (2019), Ehikioya et al. (2020), and Aziz and N'Zué (2020) just to cite a few. Greenidge et al. (2012) studied the threshold effects between public debt and economic growth in the Caribbean. Their study confirmed the existence of a debt to gross domestic product (GDP) threshold of 55-56%. They also found that debt dynamics began to change well before this threshold was reached. Specifically, at debt levels below 30% of GDP, increases in the debt-to-GDP ratio are associated with faster economic growth. However, as debt rises above 30%, the effects on economic growth decline rapidly and at debt levels reaching 55-56% of GDP, the impacts on growth shift from positive to negative. Thus, beyond this threshold, debt becomes a drag on growth. Mencinger et al. (2014) studied the direct effect of higher debt on economic growth for 25 EU countries. Their sample of EU countries was divided into subgroups to distinguish between "old" member states, over the period 1980-2010, and "new" member states, covering the period from 1995 to 2010. Using a panel estimation method, they confirmed the existence of a non-linear relationship between the ratios of government debt to GDP per capita. They also found that the threshold for the debt-to-GDP ratio should be approximately between 80 and 90% for the "old" member states. Yet, for the "new" member states, the threshold for the debt-to-GDP ratio is lower and was between 53 and 54%.

Faye and Thiam (2015) use a nested generation's model to study the effect of public debt on consumption, GDP, savings, budgetary revenues, investment, and capital dynamics in Senegal. The results show that a 10% increase in public debt positively affects macroeconomic variables but worsens the current account deficit. To be effective, a public debt of at least 65% of GDP should be integrated into the capital accumulation process. A 10% increase in external debt has a positive impact on macroeconomic variables, but worsens the current account deficit. A 10% increase in domestic debt leads to a recession. An increase in debt-financed public spending leads to an increase in the public debt bequeathed to future generations by 15% and an increase in future consumption by about 2%. As for Wade (2015), the author estimated the impact of total public debt-to-GDP ratio on the growth rate of GDP per capita with a PSTR (Panel Smooth Transition Regression Model) and the Generalized Method of Moments (GMM) method. The study covered the eight (8) WAEMU countries over a period ranging from 1980 to 2011. The results obtained with the GMM method indicated that the optimal public debt threshold stood at 48.8% of GDP, while for the PSTR the threshold stood at 49.8% of GDP. Omotosho et al. (2016) investigated the existence of threshold effects in the relationship between public debt and economic growth in Nigeria using quarterly data.

They found empirical support for an inverted U-shape relationship between public debt types and economic growth. For total public debt as percentage of GDP, the threshold level stood at 73.70%. Adeniyi et al. (2018) investigated the possible role of domestic investment in the non-linear relation between external debt and economic growth in Nigeria over the period from 1981 to 2015 using threshold regression analysis. They found that the impact of external debt on economic growth is sensitive to the measures of external debt used, and whether or not the role of domestic investment is accounted for. Accounting for the role of domestic investment in the non-linear relation between external debt and economic growth reduces the optimal debt carrying capacity of the country. Moreover, the study provided support to the crowding-out effect of excessive external debt servicing. They therefore suggested that the Nigerian government internalizes a maximum ceiling of 6.81% as the share of external debt stock in gross national income (GNI) so as to enjoy the resulting growth benefits.

N'Zué (2018), using the model of Patillo et al. (2002) with data ranging from 1970 to 2015, studied the link between external debt and growth in Côte d'Ivoire. He estimated a critical threshold of 42.9% beyond which external debt accumulation will have a negative impact on growth. Mary et al. (2019) examined the optimal point beyond which government debt impairs economic performance in Nigeria. Data from the Central Bank of Nigeria Statistical Bulletin from 1986 to 2017 were used. Dynamic Ordinary Least Square estimation method was applied. They found a significant relationship between government debt and Nigeria's economic performance. Government debt is growth-enhancing at low levels but growth-retarding at a high level with the optimal government debt estimated as 9.98% of the gross domestic product implying that borrowing beyond such a limit becomes growth retarding in the economy. Thus, government should focus on other sources of revenue to fund its budget deficits to decrease the debt burden. Aziz and N'Zué (2020), revisited the above study by using an ARDL method with data covering the period ranging from 1980 to 2018, they estimated the external debt to GDP ratio threshold to be at 59.53%. The rate beyond which debt accumulation will have a negative impact on growth.

MATERIALS AND METHODS

Following previous studies that is, Patillo et al. (2002) and N'Zué (2018), we use a Solow (1956) type production function. The theoretical framework is the neoclassical growth theory where output (Y) is a function of labor (L) and capital (K). The production function is represented below:

$$Y_t = Af(L_t, K_t) \quad (1)$$

Where, L and K are as defined above and A is a parameter that captures the effect of other factors on output. By definition, A

measures total factor productivity (*TFP*). It is through *A* that the effect of government debt on economic growth is captured.

Using the above formulation and referring to recent work, additional variables (control variables) are included in the model to help explain the output. The variable of interest "government external debt" as a percentage of GDP enters the model in both linear and quadratic terms. The quadratic term allows us to determine the threshold if it exists. The other control variables are inflation, gross fixed capital formation, working age population (used as a proxy for labor). It is important to remember that the control variables are included in the initial model (equation 1) to improve the specification of the model and to determine the effects of these other variables on the dependent variable. Equation 1 is rewritten as follows:

$$Y_t = \alpha_t + \beta X_t + \gamma Det_t + \varepsilon_t \quad (2)$$

Where, Y_t is the dependent variable; X_t is a set of control variables, (Det_t) is our variable of interest, α_t , β , γ are parameters to be estimated. t is the time period and ranges from 1980 to 2018. ε_t is the error term. As mentioned earlier, the control variables of the model include the capital variable which is represented by gross fixed capital formation ($Fcbf_t$). It measures the impact of physical capital in the production process; it is expected to be positive; the budget deficit ($Solde_t$) is included to capture the impact of fiscal policies on growth and its coefficient expected also to be positive; the trade openness indicator ($open_t$) is defined as the sum of exports-to-GDP ratio and imports-to-GDP ratio. It is introduced to capture the extent to which knowledge/technology transfer through trade impacts influences GDP. The coefficient associated with this variable is also expected to be positive. The other variables are the ratio of external debt-to-GDP ratio ($Debt_t$), terms of trade ($Term_t$), and working age population as a percentage of total population (pop_t). The linear term of external debt-to-GDP ratio is expected to have a positive coefficient while that of its quadratic term is expected to be negative.

The terms of trade variable was obtained by taking the ratio of the unit value index of exports to the unit value index of imports. It is expected to be associated with an ambiguous sign. The positive coefficient will be an indication that the terms of trade have been beneficial to the country's economy while a negative sign will indicate the extent to which the terms of trade have been detrimental to the country's economy. The working-age population is the population aged 15-64, as a percentage of the total population. It should be noted that the variables are transformed using their logarithm (ln). The data used range from 1990 to 2018 and are mainly obtained from the World Bank's World Development Indicators (WDI)², the National Directorate of Planning and Prospecting (Revised Macroeconomic Framework). Given the time-series nature of the available data, it is important to assess their time series characteristics. This includes testing for stationarity, as regressing a non-stationary variable on other non-stationary variables can lead to spurious regression. Once the assessment of the time series characteristics of the variables is completed, the next step will be to investigate the long-term dynamics of the model, which will be done by conducting co-integration tests to assess the extent to which the variables in the model move together or not in the long run. This will be done by using an *ARDL* approach and the bounds test proposed by Pesaran et al. (2001). To undertake the bounds test, it is important to reformulate the initial model to take into account both short and long term dynamics). The generalized *ARDL* (p, q) model is as follows:

$$Y_t = \alpha_t + \sum_{i=0}^p \delta_i Y_{t-i} + \sum_{i=0}^q \beta_i X_{t-i} + \varepsilon_t \quad (3)$$

Where, Y_t = Endogenous variable; X_t = Explanatory variables; α =

Constant; δ and β =parameters to be estimated; p and q = Optimal lag orders (The lags p and q are determined by minimizing the Akaike criterion (AIC) ; ε_t = Error term.

$$\begin{aligned} \Delta \ln Pibh_t = & \alpha_t + \delta_1 \ln pibh_{t-1} + \delta_2 \ln Term_{t-1} + \delta_3 \ln Fcbf_{t-1} + \\ & \delta_4 \ln Ouv_{t-1} + \delta_5 Solde_{t-1} + \delta_6 \ln Debt_{t-1} + \delta_7 \ln Debtsq_{t-1} + \\ & \delta_8 \ln Inflat_{t-1} + \delta_9 \ln pop_{t-1} + \sum_{i=0}^q \beta_{1i} \Delta \ln pibh_{t-i} + \\ & \sum_{i=0}^q \beta_{2i} \Delta \ln Term_{t-i} + \sum_{i=0}^q \beta_{3i} \Delta \ln Fcbf_{t-i} + \\ & \sum_{i=0}^q \beta_{4i} \Delta \ln Ouv_{t-i} + \sum_{i=0}^q \beta_{5i} \Delta Solde_{t-i} + \sum_{i=0}^q \beta_{6i} \Delta \ln Debt_{t-i} + \\ & \sum_{i=0}^q \beta_{7i} \Delta \ln Debtsq_{t-i} + \sum_{i=0}^q \beta_{8i} \Delta \ln Inflat_{t-i} + \\ & \sum_{i=0}^q \beta_{9i} \Delta \ln pop_{t-i} + \varepsilon_t \end{aligned} \quad (4)$$

The coefficients β_1 to β_9 represent the short-run dynamics while the coefficients δ_1 to δ_9 represent the long-run dynamics of the model. The bounds test for co-integration is equivalent to testing the following hypotheses for the above equation:

$$\begin{aligned} H_0: & \delta_1 = \delta_2 = \delta_3 = \delta_4 = \delta_5 = \delta_6 = \delta_7 = \delta_8 = \delta_9 = 0 \\ H_1: & \delta_1 \neq \delta_2 \neq \delta_3 \neq \delta_4 \neq \delta_5 \neq \delta_6 \neq \delta_7 \neq \delta_8 \neq \delta_9 = 0 \end{aligned} \quad (5)$$

This test is currently a test of the hypothesis of no co-integration among the variables (H_0) against the alternative that the variables are co-integrated (H_1) as shown above. The statistic underlying this test is the F-statistic (Pesaran et al., 1999). The calculated value of the Fisher statistic is used to decide whether or not the series are co-integrated. The asymptotic distribution of this test is non-standardized under the null hypothesis of non-co-integration between the variables. However, Pesaran et al. (2001) have provided asymptotic critical value bounds for all classifications of the regressors in $I(1)$ and or $I(0)$. Thus, if the calculated F-statistic is above the upper bound, the null hypothesis is rejected and it is concluded that there is co-integration between the variables. If it is lower than the lower limit, then the null hypothesis cannot be rejected. The conclusion is that there is no co-integration relationship between the variables. If the F-statistics are between the two limits, the test is not conclusive. When co-integration exists, the model can be rewritten to show the speed of adjustment after an exogenous shock.

After estimating equation 4, the threshold for the government external debt-to-GDP ratio is obtained by taking the first derivative of the dependent variable with respect to the debt variable and setting it to zero (equation 7).

$$\frac{\partial \ln pibh_t}{\partial \ln Debt_{t-1}} = \beta_6 + 2\beta_7 \ln Debtsq_{t-1} \quad (6)$$

$$\rightarrow \frac{\partial \ln pibh_t}{\partial \ln Debt_{t-1}} = 0 \quad (7)$$

$$\rightarrow \beta_6 + 2\beta_7 \ln Debtsq_{t-1} = 0 \quad (8)$$

$$\rightarrow \ln \widehat{Debtsq}_{t-1} = \frac{-\beta_6}{2\beta_7} \text{ avec } \beta_7 < 0 \quad (9)$$

$$\rightarrow \widehat{Debtsq}_{t-1} = e^{(-\beta_6/2\beta_7)} \quad (10)$$

Equation 10 is used to calculate the estimated threshold level of the external debt-to -GDP ratio.

RESULTS AND DISCUSSION

This section presents and discusses the empirical results. It begins with the descriptive statistics presented in Table 1. It can be observed that, on average, gross fixed capital

² World Bank (2020)

Table 1. Results of descriptive statistics for the variables of interest.

Variable	Obs	Mean	Std. Dev.	Min	Max
$pibh_t$	29	563.51	175.90	322.41	983.30
$dette_t$	29	69.54	31.81	20.40	115.53
$fbcf_t$	29	22.55	7.49	14.29	54.30
pop_t	29	63.44	0.85	61.48	64.31

Author's calculation.

Table 2. Results of the unit root tests using the "Augmented Dickey Fuller" and "Philip Perron" tests.

Variable	ADF		PP		Decision
	Level	1st difference	Level	1st difference	
$\ln pibh_t$	-0.987(-3.588)		0.367(-2.992)	-3.854(-3.592)	I(1)
$\ln dette_t$	-1.793(-3.588)	-4.753(-3.592)	-0.310(-2.992)	-5.399(-3.592)	I(1)
$\ln fbcf_t$	-4.286(-3.588)		-3.150(-2.992)		I(0)
$\ln detteq_t$	-1.732(-3.588)	-4.440(-3.596)	0.018(-2.992)	-5.253(-3.592)	I(1)
$\ln pop_t$	-0.879(-3.588)		5.613(-2.992)	-4.620(-3.592)	I(1)

Author's calculation.

Table 3. Test of co-integration between the variables of interest for *ARDL* (1,5).

H0: No long run relationship		
F-stat	F=10.42	
K=4	I (0)	I (1)
Critical value at 10%	2.45	3.52
Critical value at 5%	2.86	4.01
Critical value at 1%	3.74	5.06

Accept H_0 if $F_{stat} < \text{Critical value for } I(0)$; Reject H_0 if $F_{stat} > \text{Critical value for } I(1)$
 Source: Author's calculation.

formation is very high in Guinea. Indeed, it was 22.55% and above the ECOWAS regional threshold of 20%. On average, the minimum debt-to-GDP ratio was 20.41% and was recorded in 2012 after the country benefited from the Heavily Indebted Poor Countries (HIPC) initiative. The highest debt-to-GDP ratio (115.54%) was recorded in 2005. The time series characteristics of the variables were analyzed (Table 2). The results show that with the exception of the variable gross fixed capital formation which is integrated of order 0, that is $I(0)$, all other variables are integrated of order 1, that is $I(1)$. The above results, show a mixture of $I(0)$ and $I(1)$ variables confirming the use of the *ARDL* (p, q) approach. The results of the bounds tests are presented in Table 3. The F-statistic is compared to the critical bound test values

tabulated by Pesaran et al. (2001) without restriction on the constant and trend. The null hypothesis of the test is that there is no co-integrating relationship versus the alternative hypothesis of a co-integrating relationship. The value of the F-statistic calculated is 10.42. It is greater than all of the critical values considered, namely 1%, 5% and 10%. The null hypothesis of no co-integration cannot be accepted. It is therefore concluded that there is a co-integration relationship between the variables, which means that they move together in the long run.

With the above result, we proceed to estimate the short and long term dynamics. The results are presented in Table 4. From Table 4, we found that in the long run, debt is positively related to growth and the coefficient

Table 4. Results of the estimated *ARDL(1,5)* model (1).

Variable	Dependent variable : Per capita gross domestic product	
	Coefficients	Probability Value
Long run dynamics		
\ln^{pibh}_t (ADJ)	-0.976*	0.000
\ln^{dette}_t	2.907*	0.000
\ln^{fbcf}_t	0.503*	0.002
\ln^{detteq}_t	-0.450*	0.000
\ln^{pop}_t	16.195**	0.026
Short run dynamics		
\ln^{pibh}_t	0.243	0.105
\ln^{dette}_t	0.121	0.907
\ln^{fbcf}_t	0.482*	0.002
\ln^{pop}_t	-39.618	0.183
C		
R-square \rightarrow 0.881	Adjusted R-square \rightarrow 0.780	
Autocorrelation test (Breusch-Godfrey)		
F-stat= 16.02	P-Value F= 0.034	
Heteroskedasticity test (White)		
F-stat= 25.80	P-Value F= 0.081	
Normality test (Jarque Bera) \rightarrow 0.442		
Ramsey test		
F-stat= 2.85	P-Value F= 0.061	

Asterisks, *, **, *** indicate significance at 1, 5 and 10%, respectively.

Source: Author, based on data from the Revised Macroeconomic Framework of Guinea and WDI (2018).

associated with the squared variable is negative. This indicates that in the long run there is an optimal level of debt beyond which it will have a negative effect on growth. The error correction term is negative and significant, confirming the co-integration relationship between the variables. Furthermore, in the long term dynamics, we observe that all the variables are significant at 5%. With the above results (long-term dynamics) and using equation 10, it is possible to estimate the optimal level of indebtedness beyond which an increase in the external debt-to-GDP ratio will have a negative effect on the country's economy. Indeed, replacing the estimated parameters in equation 10 enables us to obtain the estimated optimal level of external debt-to-GDP ratio. It stood at 25.7%. Thus, beyond this point, an increase of the external debt-to-GDP ratio resulting from an increase in external debt will have a negative impact on the country's economic performance. It was also found that in the case of Guinea, the external debt-to-GDP ratio has no effect on the country's economic performance.

Conclusion

The objective of this study was to determine the threshold of Guinea's external public debt-to-GDP ratio and compare it to the current evolution of public debt. Specifically, the study sought to determine the impact of the external public debt-to-GDP ratio on the country's economic performance; and to determine the threshold level of the external public debt-to-GDP ratio beyond which economic performance would be affected. We used an *ARDL* approach. We found that the variables considered in this study are co-integrated. That is, they move together in the long run. The *ARDL(1,5)* model estimated enabled us to have the following results: In the long run, external debt-to-GDP ratio has a positive impact on the country's economic performance; there is a threshold level of external debt-to-GDP ratio beyond which its impact on economic performance is negative. That threshold level stood at 25.27%. Considering the current level of the external debt-to-GDP ratio, which

stood at 21.7%, it is clear that the country still has room to more borrowing, however we should call for caution as this level is not too far from the threshold.

This result is in line with previous studies (Omotosho et al. (2016), Adeniyi et al. (2018), N'Zué (2018), Mary et al. (2019)) that found a threshold level for external debt: in the long run, a 1% increase in the stock of external public debt-to-GDP ratio could lead to a 2.9% increase in per capita GDP; a 1% increase in investment as a percentage of GDP will lead to a 0.5% increase in per capita GDP; and a 1% increase in the working age population as a percentage of total population will lead to a 16.1% increase of per capita GDP. The results are also in line with Ehikioya et al. (2020) who found long run equilibrium between external debt and economic growth. In the short run, we obtained the following results: external debt-to-GDP ratio is positively related to economic growth but it is not significant. Investment as a percentage of GDP has a positive and significant impact on GDP per capita thus a 1% increase in the investment variable leads to a 0.48% increase in GDP per capita and the working age population has a negative impact on GDP per capita.

Recommendations

Given that investment has a positive impact on economic growth, it is necessary for the country's authorities to encourage the development of public investment policies that promote the private sector. Additional research is needed to undertake a thorough assessment of the utilization of the resources borrowed. Also, they should develop the skills of the working age population to boost further its impact on the country's economic performance.

CONFLICT OF INTERESTS

The authors have not declared any conflict of interest.

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